

## Expectations

"weighted average"

$$EX = \sum_a a p(a) = \sum_a a P\{X=a\}$$

Ex

$X = \#$  of H's in 3 independent tosses of a fair coin.

$$p(0) = p(3) = 1/8$$

$$p(1) = p(2) = 3/8$$

$$\} \Rightarrow EX = 1.5$$

Ex

$X$  is  $\text{Bin}(n, p)$

$$EX = \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n \frac{n!}{(n-k)! (k-1)!} p^k (1-p)^{n-k}$$

$$= np \sum_{k=1}^n \frac{(n-1)!}{[(n-1)-(k-1)]! (k-1)!} p^{k-1} (1-p)^{(n-1)-(k-1)}$$

$$= np \sum_{j=0}^{n-1} P\{\text{Bin}(n-1, p) = j\}$$

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$$= np$$

Ex.  $X = \text{Poisson}(\lambda) \Rightarrow$

$$EX = \sum_{k=0}^{\infty} k \frac{e^{-\lambda} \lambda^k}{k!} = \lambda e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} = \lambda$$

Ex  $X =$  a neg. binomial r.v.  $(r/p)$ ; i.e.,

$$p(k) = \binom{k-1}{r-1} p^r (1-p)^{k-r} \quad r=1,2,3,\dots$$

$$\begin{aligned} E(X) &= \sum_{k=r}^{\infty} k \binom{k-1}{r-1} p^r (1-p)^{k-r} \\ &= \sum_{k=r}^{\infty} \frac{k!}{(r-1)!(k-r)!} p^r (1-p)^{k-r} \end{aligned}$$

$$= \frac{r}{p} \sum_{k=r}^{\infty} \frac{k!}{r!(k-r)!} p^{r+1} (1-p)^{(k+1)-(r+1)}$$

$$= \frac{r}{p} \sum_{j=r+1}^{\infty} \binom{j-1}{(r+1)-1} p^{r+1} (1-p)^{j-(r+1)}$$

$$= r/p.$$

Ex (Indicator Function)  $1_A = \begin{cases} 1, & \text{if } A \\ 0, & \text{if } A^c \end{cases}$

$$P(1_A = 1) = P(A) \quad \rightarrow \quad E(1_A) = P(A).$$

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Functions of a r.v.

$Y = X^2$  is a r.v.  $P_Y(k) = ?$ ,  $EY = ?$ , etc.

Ex -  $X = \begin{cases} 1 \\ -1 \end{cases}$  wp  $\frac{1}{2}$  each ( $EX = 0$ )

$Y = X^2 = 1$  wp 1. ( $EY = 1$ )

Ex -  $X = 0, \pm 1, \pm 2$  wp  $\frac{1}{5}$  each.  $Y = X^2$

$Y = \begin{cases} 0 & 1/5 \\ 1 & 2/5 \\ 4 & 2/5 \end{cases} \Rightarrow EY = EX^2 = \frac{10}{5} = 2.$

So  $E(X^2) \neq (EX)^2$ ; in fact there is no obvious relation between  $E[g(X)]$  and  $EX$ ... ever!

To compute  $E[g(X)]$ , the 1<sup>st</sup> method is to find the prob. dist. of  $Y = g(X)$ , and proceed. This is too arduous. Alternatively, use

Theorem If  $X$  has possible values  $x_1, x_2, \dots$  with prob.  $p(x_1), p(x_2), \dots$  then

$$E[g(X)] = \sum_{i=1}^{\infty} g(x_i) p(x_i).$$

Proof  $E[g(X)] = \sum_{i \geq 1} \cancel{g(x_i)} y_i \cdot P\{g(X) = y_i\}$  ( $y_i$  = <sup>different</sup> values of  $g(x_i)$ )

$$= \sum_{i \geq 1} y_i \sum_{\substack{X=x_j \\ j: g(x_j)=y_i}} P\{X=x_j\} = \sum_j \sum_{i: y_i=g(x_j)} y_i p(x_j)$$

$$= \sum_j \sum_{i: y_i=g(x_j)} g(x_j) p(x_j)$$

$$= \sum_j g(x_j) p(x_j) \quad \#$$

Cor If  $a, b \in \mathbb{R}$  then  $E[aX+b] = aE(X) + b$ .

Proof  $E(aX+b) = \sum_{i \geq 1} (ax_i + b) p(x_i)$

$$= a \sum_{i \geq 1} x_i p(x_i) + b \sum_{i \geq 1} p(x_i)$$

$$= aEX + b \quad \star$$

Ex  $X: \text{Bin}(n, p)$

$$EX^2 = \sum_{k=0}^n k^2 \binom{n}{k} p^k (1-p)^{n-k}$$

$$= \sum_{k=0}^n k^2 \frac{n!}{(n-k)! k!} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n k \frac{n}{(n-k)! (k-1)!} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n (k-1) \frac{n!}{(n-k)! (k-1)!} p^k (1-p)^{n-k}$$

$$+ \sum_{k=1}^n \frac{n!}{(n-k)! (k-1)!} p^k (1-p)^{n-k}$$

~~$$= \sum_{k=2}^n \frac{n!}{(n-k)! (k-2)!} p^{k-2} (1-p)^{(n-2)-(k-2)}$$~~

$$T_1 = \sum_{k=2}^n \frac{(n-2)!}{((n-2)-(k-2))! (k-2)!} p^{k-2} (1-p)^{(n-2)-(k-2)} \cdot \frac{(n-1) n p^2}{(k-2)}$$

$$= \sum_{j=0}^{n-2} \binom{n-2}{j} p^j (1-p)^{n-2-j} \frac{n(n-1)}{p} = \frac{np^2 (n-1)}{p}$$

$$T_2 = \sum_{k=1}^n \frac{n!}{(n-k-1)!(k-1)!} p^{k-1} (1-p)^{(n-1)-(k-1)} = p$$

$$= np \sum_{k=1}^n \binom{n-1}{n-1-k-1} p^{k-1} (1-p)^{(n-1)-(k-1)}$$

$$= np \sum_{j=0}^{n-1} \binom{n-1}{j} p^j (1-p)^{n-1-j} = np.$$

$$\therefore EX^2 = n(n-1)p^2 + np = n^2p^2 - np^2 + np$$

$$= n^2p^2 + np[1-p] = (EX)^2 + np(1-p).$$